Diffusion Prior for Online Decision Making: A Case Study of Thompson Sampling

Yu-Guan Hsieh∗  Shiva Kasiviswanathan
Université Grenoble Alpes  Amazon
yu-guan.hsieh@univ-grenoble-alpes.fr  kasivisw@amazon.com

Branislav Kveton  Patrick Bloebaum
AWS AI Labs  Amazon
bkveton@amazon.com  bloebp@amazon.com

Abstract

In this work, we investigate the possibility of using denoising diffusion models to learn priors for online decision making problems. Our special focus is on the meta-learning for bandit framework, with the goal of learning a strategy that performs well across bandit tasks of a same class. To this end, we train a diffusion model that learns the underlying task distribution and combine Thompson sampling with the learned prior to deal with new task at test time. Our posterior sampling algorithm is designed to carefully balance between the learned prior and the noisy observations that come from the learner’s interaction with the environment. Preliminary experiments clearly demonstrate the potential of the considered approach.

1 Introduction

Uncertainty quantification is an integral part of online decision making and forms the basis of various online algorithms that trade-off exploration against exploitation. Among these methods, Bayesian approaches allow us to quantify the uncertainty using probability distributions, with the help of the powerful tools of Bayesian inference. Nonetheless, their performance is known to be sensitive to the choice of prior.

For concreteness, let us consider the problem of stochastic multi-armed bandits (MABs) [2, 10], in which a learner repeatedly pulls one of the $K$ arms from a given set $A = \{1, \ldots, K\}$ and receives rewards that depend on the learner’s choices. More precisely, when arm $a$ is pulled at round $t$, the learner receives reward $r_t$ drawn from an arm-dependent distribution $\mathcal{P}^a$. The goal of the learner is either to i) accumulate the highest possible reward over time (a.k.a. regret-minimization) or to ii) find the arm with the highest expected reward within a prescribed number of rounds (a.k.a. best-arm identification).

For both purposes, we need to have a reasonable estimate of the arms’ mean rewards $\mu^a = \mathbb{E}_{r \sim \mathcal{P}^a}[r^a]$. In general, this would require us to pull each arm a certain number of times, which becomes inefficient when $K$ is large. While the no-free-lunch principle prevents us from improving upon this bottleneck in general situations, it is worth noticing that the bandit instances (referred as tasks hereinafter) that we encounter in most practical problems are far from arbitrary. To name a few examples, in recommendation systems, each task corresponds to a user with certain underlying preferences that affect how much they like each item; in online shortest path routing, we operate in real-world networks that feature specific characteristics. In this regard, introducing such inductive bias to the learning algorithm would be beneficial. In Bayesian models, this can be expressed through

∗Work done during internship at Amazon.

NeurIPS 2022 Workshop on Score-Based Methods.
the choice of the prior distribution. Moreover, as suggested by the meta-learning paradigm, the prior itself can also be learned from data, which often leads to superior performance \[7, 13\].

**Our contributions.** This work tackles the problem of meta-learning a prior for bandits \[1, 3, 11\], with Thompson sampling used as the base algorithm in each task. In order to approximate the complex priors that arise in practice, we build upon the powerful tools of deep generative modeling, whose recent progress have enabled impressive results in various areas ranging from image generation \[15\] to protein design \[21\]. Specifically, we propose to meta-learn the prior using denoising diffusion models \[6, 17\], and develop an algorithm to perform Thompson sampling under the learned prior. The designed algorithm strikes a delicate balance between the learned prior and bandit observations, bearing in mind the importance of having an accurate uncertainty estimate. Through synthetic experiments, we demonstrate the benefit of the considered approach against several baseline methods.

## 2 Preliminaries and Problem Description

In this section, we briefly review denoising diffusion models and introduce our meta-learning framework over a class of bandit tasks. The associated pseudo-codes can be found in Appendix A.

### 2.1 Denoising Diffusion Probabilistic Model

First introduced by Sohl-Dickstein et al. \[17\] and recently popularized by Ho et al. \[6\] and Song and Ermon \[18\], denoising diffusion models (or the closely related score-based models) have been shown to achieve state-of-the-art performance in various data generation tasks. Numerous variants of these models have been proposed. Below, we mainly adopt the notations and formulation of Ho et al. \[6\], with minimal modifications to adapt it to our purpose.

Intuitively speaking, diffusion models learn to approximate a distribution \(Q_0\) by training a series of denoisers with samples drawn from this distribution. Writing \(q\) for the probability density function (assume everything is Lebesgue measurable for simplicity) and \(X_0\) for the associated random variable, we define the forward diffusion process with respect to a sequence of scale factors \((\alpha_\ell) \in (0, 1)^L\) by

\[
q(x_{1:L} | x_0) = \prod_{\ell=0}^{L-1} q(x_{\ell+1} | x_\ell), \quad q(X_{\ell+1} | x_\ell) = N(X_{\ell+1}; \sqrt{\alpha_\ell} x_\ell, (1 - \alpha_\ell) I).
\]

The first equality suggests that the forward process is Markovian, while the second equality implies that the transition kernel is Gaussian. Further denoting the product of the scale factors by \(\tilde{\alpha}_\ell = \prod_{\ell=1}^{\ell} \alpha_\ell\), we then have \(q(X_\ell | x_0) = N(X_\ell; \sqrt{\tilde{\alpha}_\ell} x_0, (1 - \tilde{\alpha}_\ell) I)\).

The sequence \((\alpha_\ell) \in (0, 1)^L\) is chosen to be decreasing and such that \(\tilde{\alpha}_L \approx 0\). We thus expect \(q(X_L) \approx N(0, 1)\). A denoising diffusion model learns to reverse the diffusion process with a distribution \(P_\theta\) over random variables \(X_{0:L}\), in the hope that the marginal distribution \(P_\theta(X_0)\) is a good approximation of \(Q_0\). This is achieved by setting \(p_\theta(X_\ell | X_{\ell+1}) = N(0, 1)\), enforcing the learned reverse process to be Markovian, and modeling \(p_\theta(X_{\ell} | x_{\ell+1})\) as a Gaussian parameterized by

\[
p_\theta(X_\ell | x_{\ell+1}) = q(X_\ell | x_{\ell+1}, X_0 = h_\theta(x_{\ell+1}, \ell + 1)) \propto q(x_{\ell+1} | X_\ell) q(X_\ell | X_0 = \hat{x}_0).
\]

In the above \(h_\theta\) is the learned denoiser and \(h_\theta(x_{\ell+1}, \ell + 1)\) is the predicted clean sample.\(^2\)

### 2.2 Meta-Learning of Bandit Tasks

We consider a meta-learning for bandits framework in which the bandit tasks are drawn from an underlying distribution \(\mathcal{T}\). Unlike \[3, 11\], we focus on the a more standard scenario which features a meta-train and a meta-test phase. The goal is to learn a good prior using the training set such that we can perform better once the algorithm is deployed with the learned prior in the test phase. Mathematically, we can characterize the performance of the algorithm with the so-called *transfer regret* \[3\], defined by \(\text{Reg}_T(\pi) = E_{B \sim T} \text{Reg}_T(\pi, B)\) where \(\pi\) is the algorithm that uses the learned prior and \(\text{Reg}_T(\pi, B)\) is the regret incurred by the algorithm within task \(B\) (see Appendix C).

\(^2\)To obtain \(h_\theta\) we typically train a neural network with a U-Net architecture. In \[6\], this network is trained to output the predicted noise \(\hat{z}_\ell = (x_\ell - \sqrt{\alpha_\ell} h_\theta(x_\ell, \ell))/\sqrt{1 - \alpha_\ell}\).
Throughout the work, we assume that the noise in the rewards are Gaussian with known variance $\sigma^2$. The learner thus only needs to learn the vector of the mean rewards $\mu = (\mu^a)_{a \in A}$. Accordingly, the prior is defined as a distribution in $\mathbb{R}^K$.

3 Algorithm

In this section, we describe our backbone algorithm for Thompson sampling with diffusion prior.

3.1 Diffusion Model as Prior

Our training data for the diffusion prior can have different forms. For example, these data may be composed of the interaction history with tasks using some bandit algorithms. How to leverage such confounded and noisy data to learn a diffusion model is a challenging problem. In this short paper, we simply assume that the mean vectors of these tasks are known, from which we can apply the standard diffusion model training process.

Once the diffusion model is learned, we may use it directly as a prior in any downstream tasks. Nonetheless, the original sampling process of the diffusion model as described in [6] falls short in providing a good estimate of the variances that reflects the right level of uncertainty. To address this problem, we propose to further calibrate the variances of the reverse process with a separate calibration set $\mathcal{V}$. For this, we write

$$p_0(X_{\ell} | x_{\ell+1}) = \int q(X_{\ell} | x_{\ell+1}, x_0) p_0(x_0 | x_{\ell+1}) \, dx_0.$$  \hfill (2)

In the above, $p_0'(X_0 | x_{\ell+1})$ is a Gaussian distribution centered at $\tilde{x}_0 = h_0(x_{\ell+1}, \ell + 1)$. This is different from (1) where by analogy we may interpret $p_0'(x_0 | x_{\ell+1}) dx_0$ as a Dirac. The covariance of $p_0'(X_0 | x_{\ell+1})$ is then taken as a diagonal matrix $\text{diag}(\tau_0^2)$ whose diagonal elements are the (coordinate-wise) mean squared reconstruction errors of the model on the calibration set $\mathcal{V}$. That is, we construct $\mathcal{V}_t$ containing pairs $(x_0, x_{\ell})$ with $x_0 \in \mathcal{V}$ and $x_{\ell}$ sampled from $X_{\ell} | x_0$, and set $\tau_0^2 = \sqrt{\sum_{x_{\ell}, x_0 \in \mathcal{V}_t} \| x_0 - h_0'(x_0, \ell) \|^2 / \text{card}(\mathcal{V}_t)}$. Intuitively, this adjusts how much we rely on the learned model in the upcoming tasks by taking the reconstruction error as a proxy for its quality.

3.2 Thompson Sampling with Diffusion Prior

Thompson sampling [14, 20] is one of the most popular strategies for tackling stochastic bandits due to its simplicity and generality. It takes as input a prior distribution for the parameter of interest, and samples a guess of the parameter from the posterior distribution at each round to determine which arm to pull. In the MAB setting in Section 1, the parameter of interest is the vector of mean rewards $\mu$. At each round $t$, a vector $\mu_t$ is sampled from the posterior distribution determined by the prior and the interaction history $(a_s, r_s)_{s \in \{1, \ldots, t-1\}}$. Subsequently, the learner pulls an arm $a_t \in \mathbb{arg\ max}_{a \in A} \hat{\mu}_t^a$.

In the following we design an algorithm that samples from the posterior distribution when the prior is described by a diffusion model. While an exact solution does not exist in general, we aim to find a good approximation that balances well between the meta-learned prior and noisy observations. In contrast, most existing approaches for posterior sampling with diffusion models focus on the solution of inverse problems [5, 8, 17, 19], and do not take the underlying uncertainty into account.

To proceed, the general idea of our algorithm is to guide the sample towards the observation during the sampling process. Formally, let us consider $Y_0$ as a random observation of $X_0$ with known $q(Y_0 | X_0)$. For given $y_0$, we are interested in sampling from $X_0 | y_0$. We achieve this by going through the reverse Markovian process, conditioning on $Y_0 = y_0$. In fact, to sample from $X_{\ell} | y_0$, we only need to first sample $x_{\ell+1}$ from $X_{\ell+1} | y_0$ and then sample $x_{\ell}$ from $X_{\ell} | x_{\ell+1}, y_0$. Repeating the process for $\ell = L - 1, \ldots, 0$ then gives the desired result. Below we briefly explain the initialization and the recursive steps of our method. Detailed derivation of the algorithm along with alternative approaches are discussed in Appendix B.

Sampling from $X_L | y_0$. For this part, we simply ignore $y_0$ and sample from $\mathcal{N}(0, 1)$ as before.

Sampling from $X_{\ell} | x_{\ell+1}, y_0$. For simplicity, we restrict our attention to the case of multi-armed bandit. Here, $y_0$ is the interaction history and $x_0 = \mu$ is the vector of mean rewards, leading to

$$q(y_0 | x_0) \propto \prod_{s=1}^{t} q(r_s | \mu, a_s) = \prod_{s=1}^{t} \mathcal{N}(r_s; \mu^{a_s}, \sigma^2) \quad \text{[as a function of } x_0].$$  \hfill (3)
We distinguish between two situations.

- Arm $a$ has never been pulled in the first $t$ rounds: We sample the corresponding coordinate $x_{t}^{a}$ from the Gaussian distribution $\tilde{q}(x_{t}^{a} \mid x_{t+1}, y_{0}) = p_{\theta}(x_{t} \mid x_{t+1})$ introduced in (2).
- Arm $a$ has been pulled in the first $t$ rounds: We denote by $\hat{\mu}_{t}^{a} = \sum_{s=1}^{t} r_{s} \mathbb{1}\{a_{s} = a\}/N_{t}^{a}$ the empirical mean and $\sigma_{t}^{a} = \sigma/\sqrt{N_{t}^{a}}$ the adjusted standard deviation, where $N_{t}^{a}$ is the number of times that arm $a$ has been pulled up to time $t$ (included). We also write $\tilde{z}_{t+1}$ for the noise predicted by the denoiser from $x_{t+1}$. Then, with $\rho_{t} = \bar{\alpha}_{t+1} (1 - \bar{\alpha}_{t}) / (\bar{\alpha}_{t}(1 - \bar{\alpha}_{t+1}))$, we sample from the Gaussian distribution $\tilde{q}(x_{t}^{a} \mid x_{t+1}, y_{0})$ satisfying that

$$
\tilde{q}(x_{t}^{a} \mid x_{t+1}, y_{0}) \propto p_{\theta}(x_{t}^{a} \mid x_{t+1}) N \left( x_{t}^{a}; \sqrt{\alpha_{t}\hat{\mu}_{t}^{a}} + \sqrt{1 - \alpha_{t}} \bar{\alpha}_{t} \tilde{z}_{t+1}, \bar{\alpha}_{t}(\sigma_{t}^{a})^{2} + \rho_{t}(\tau_{t+1})^{2} \right).
$$

In words, we sample $x_{t}$ from a Gaussian whose mean is a weighted average of the mean of $p_{\theta}(x_{t}^{a} \mid x_{t+1})$ and a noisy version of the empirical mean $\hat{\mu}_{t}^{a}$ (only defined for those arms that have been pulled). Importantly, the noisy version of the observation is computed with the predicted noise, which distinguishes our work from existing ones such as [19].

4 Numerical Experiments

In this section, we illustrate the benefit of using diffusion prior through two synthetic experiments. Missing details can be found in Appendix C.

**Problem construction.** For the two problems, we fix $\sigma = 0.1$ and construct the means as follows

1. **Labeled Arms Problem.** Let $K = 500$. In this problem, each arm is associated to 7 of the 50 labels. For each bandit task, we randomly pick 7 labels; the expected reward of an arm is positively correlated to the number of labels that fall into the intersection of the arm’s labels and the task’s labels. This is a simplest model for modeling user preferences over a list of items.

2. **Popular and Niche Problem.** Let $K = 200$. In this problem, the arms are separated into 40 groups, each of size 5. Among these, 20 groups of arms tend to have high expected rewards but these arms are never the optimal ones. The other 20 groups of arms have lower expected rewards in general but contain the optimal arm. This represents a situation where Gaussian prior can lead to poor performance due to its inability to model multi-modal distributions.

**Training and baselines.** In terms of algorithms, we compare our method, DiffTS, with UCB and Thompson sampling with Gaussian prior using either diagonal or full covariance matrix (GTS-diag and GTS-full). For each algorithm, we test two configurations: either the algorithm is given the true labels or an underspecified standard deviation $\sigma' = 0.1$ or an underspecified standard deviation $\sigma' = 0.05$.

To train the diffusion model, we use a training set of 5000 samples and a calibration set of 1000 samples. The 6000 samples put together are also used to compute the means and the covariances used by the Gaussian Thompson sampling algorithms. To test the performance of the algorithms, we sample another 100 bandit tasks and run the aforementioned algorithms on these tasks.

**Results.** We report the regrets of the algorithms averaged over the 100 test tasks in Figure 1 (the top and the bottom figures correspond respectively to the Labeled Arms and the Popular and Niche problem). We verify that using a diffusion prior which describes better the task distribution indeed helps achieve smaller regret. However, the amount of improvement varies. In the Labeled Arms problem, learning the correlation between arms already lower the regret significantly, and the assumed noise standard deviation $\sigma'$ seems to have a greater impact. In the Popular and Niche problem, using a Gaussian prior with full covariance matrix however leads to poor performance, and the ability of denoising diffusion models to model complex distribution helps greatly here.

---

$^3$It holds that $x_{t+1} = \sqrt{\alpha_{t+1}} h_{\theta}(x_{t+1}, \ell + 1) + \sqrt{1 - \alpha_{t+1}} \tilde{z}_{t+1}$. 

![Fig. 1: Average regret over 100 tasks of different algorithms in the two problems. Shaded areas represent standard errors.](image-url)
References


Appendix

A Missing Pseudo Codes

Algorithm 1 Meta Learning for Bandits with Diffusion Models
1: Meta Training
2: **Input**: A set of expected means \((\mu_B)_B\) from different tasks \(B \sim \mathcal{T}\)
3: Train a diffusion model (a denoiser) \(h_\theta\) to model the distribution of the mean rewards
4: Calibration
5: **Input**: A set of expected means \((\mu_B)_B\) from different tasks \(B \sim \mathcal{T}\)
6: Compute the mean squared reconstruction error \((\tau^2_\ell)_{\ell \in \{1, \ldots, L\}}\) for the denoiser \(h_\theta\) at different noise levels to calibrate the variance
7: Meta Test / Deployment
8: For any new task \(B\), run Thompson sampling with diffusion prior using the trained model

Algorithm 2 Thompson Sampling with Diffusion Prior (DiffTS)
1: **Input**: Learned denoiser (diffusion model) \(h_\theta\), reconstruction errors \((\tau^2_\ell)_{\ell \in \{1, \ldots, L\}}\), presumed noise standard deviation \(\sigma'\)
2: for \(\ell = 1, \ldots \) do
3: Sample \(x_L \sim \mathcal{N}(0, I)\)
4: for \(\ell \in L - 1, \ldots, 0\) do
5: Predict clean sample \(\hat{x}_0 = h_\theta(x_{\ell+1}, \ell + 1)\) and associated noise \(\tilde{z}_{\ell+1}\)
6: Compute diffused observation \(\tilde{y}_\ell = \sqrt{\alpha_\ell}\hat{\mu}_{\ell-1} + \sqrt{1 - \alpha_\ell}\tilde{z}_{\ell+1}\)
7: for \(a \in \mathcal{A}\) do
8: If \(N_{\ell-1}^a = 0\), sample \(x_\ell^a \sim p_\theta(X_\ell^a | x_{\ell+1})\)
9: If \(N_{\ell-1}^a > 0\), sample \(x_\ell^a \sim q(X_\ell^a | x_{\ell+1}, y_0) \propto p_\theta(X_\ell^a | x_{\ell+1})\mathcal{N}(X_\ell^a; \hat{y}_\ell^a, \alpha_\ell((\sigma_\ell^a)^2 + \rho_\ell(\tau^2_{\ell+1}))\)
10: end for
11: end for
12: Pull arm \(a_t \in \arg\max_{a \in \mathcal{A}} x_0^a\)
13: Update number of pulls \(N_t^a\), scale std \(\sigma_t^a\), and empirical reward \(\hat{\mu}_t^a\) for \(a \in \mathcal{A}\)
14: end for

B Posterior Sampling: Algorithms and Derivation

Below we provide two approximations to sample from \(X_L | x_{\ell+1}, y_0\). The first one is easier to derive but the second one consistently achieves better performance in our experiments.\(^4\) We thus utilize the second approach in Sections 3 and 4.

1. **Approach 1**, acting on \(x_0\). We write

\[
q(x_\ell | x_{\ell+1}, y_0) = \frac{\int q(x_\ell, y_0, x_0 | x_{\ell+1}) dx_0}{q(y_0 | x_{\ell+1})} = \frac{\int q(y_0 | x_0)q(x_\ell | x_0, x_{\ell+1})q(x_0 | x_{\ell+1}) dx_0}{q(y_0 | x_{\ell+1})}.
\]

As in Section 3.1, we approximate \(q(x_0 | x_{\ell+1})\) by

\[
p'_\theta(x_0 | x_{\ell+1}) = \mathcal{N}(x_0; h_\theta(x_{\ell+1}, \ell + 1), \text{diag}(\tau^2_{\ell+1})).
\]

\(^4\)We conjecture this is because the first approach leads to results that are less consistent with the observations.
Then, if \( q(Y_0 \mid x_0) \) is also Gaussian, the integral in (4) with \( q(x_0 \mid x_{\ell+1}) \) replaced by \( p_0'(x_0 \mid x_{\ell+1}) \) can be computed in close form.

For further illustration, we focus on the case of multi-armed bandits where the relation between \( Y_0 \) the interaction history and \( x_0 = \mu \) the mean reward vector is given by (3). The approximate distribution of \( X_{\ell} \mid x_{\ell+1}, y_0 \) is thus independent across coordinates (arms) and we can compute the density individually for each arm, which we denote by \( q(x_\ell^a \mid x_{\ell+1}, y_0) \). It is computed using (4) with \( q(x_0 \mid x_{\ell+1}) \approx p_0'(x_0 \mid x_{\ell+1}) \) followed by a normalization.

When arm \( a \) has never been pulled, \( q(x_\ell^a \mid x_{\ell+1}, y_0) \) is nothing but \( p_0(x_\ell^a \mid x_{\ell+1}) \). Otherwise, \( q(x_\ell^a \mid x_{\ell+1}, y_0) \) is computed first using the fact that

\[
q(y_0^a \mid x_0^a)p_0'(x_0^a \mid x_{\ell+1}) \propto \mathcal{N} \left( x_0^a; \frac{\hat{\rho}_0^2/(\sigma_0^2)^2 \cdot h_0(x_{\ell+1}, \ell + 1)/ (\tau_{\ell+1}^a)^2 - 1/ \left( \sigma_0^2 \right)^2 + 1/ (\tau_{\ell+1}^a)^2)}{x_0^a}, \frac{1}{1/(\sigma_0^2)^2 + 1/(\tau_{\ell+1}^a)^2} \right) .
\]

Next, with

\[
q(x_\ell^a \mid x_0, x_{\ell+1}) = \mathcal{N} \left( x_\ell^a; \frac{\sqrt{\alpha_{\ell+1}(1 - \hat{\alpha}_{\ell})}}{1 - \hat{\alpha}_{\ell+1}} x_0^a + \frac{\sqrt{\alpha_{\ell+1}(1 - \hat{\alpha}_{\ell})}}{1 - \hat{\alpha}_{\ell+1}} x_{\ell+1}^a, \frac{(1 - \hat{\alpha}_{\ell}) \beta_{\ell+1}}{1 - \hat{\alpha}_{\ell+1}} \right),
\]

we see that we just need to replace \( x_0^a \) by \( x_\ell^a \) and augment the variance by \( \tau_\ell^a(\sigma_\ell^a)^2 \) when sampling \( x_\ell^a \). Compared to the case where \( y_0 \) is not given, the main difference lies in that we first change the estimated \( x_0 \) before sampling \( x_\ell \).

2. **Approach 2, acting on \( x_\ell \).** Alternatively, we may write

\[
q(x_\ell \mid x_{\ell+1}, y_0) = \frac{q(x_\ell \mid x_{\ell+1})q(y_0 \mid x_\ell, x_{\ell+1})}{q(y_0 \mid x_{\ell+1})} = \frac{q(x_\ell \mid x_{\ell+1}) \int q(y_0 \mid x_0) q(x_0 \mid x_\ell, x_{\ell+1}) \, dx_0}{q(y_0 \mid x_{\ell+1})}.
\]

To begin, we use \( p_0(x_\ell \mid x_{\ell+1}) \) to approximate \( q(x_\ell \mid x_{\ell+1}) \). Next, one natural way to tackle the integral is to use \( q(x_0 \mid x_\ell, x_{\ell+1}) = \frac{q(x_0 \mid x_\ell)}{q(x_\ell \mid x_{\ell+1})} \approx p_0'(x_0 \mid x_\ell) \). In the simplest case \( q(y_0 \mid x_0) = \mathcal{N}(y_0; x_0, \sigma I) \), and we deduce

\[
\int q(y_0 \mid x_0)p_0'(x_0 \mid x_\ell) \, dx_0 = \mathcal{N}(y_0; h_0(x_\ell, \ell), \sigma^2 I + \text{diag}(\tau_\ell^2)).
\]

Nonetheless, as the denoiser \( h_0 \) can be arbitrarily complex, this does not lead to a close form expression to sample \( x_\ell \). We may resort to Langevin dynamics sampling, i.e., taking gradient steps on \( x_\ell \) to make sure the denoised result is coherent with \( y_0 \). This is a general approach that may be of interest for arbitrary conditional distribution \( Y_0 \mid X_0 \), but we have simpler solution when the noise is Gaussian. In fact, it holds that

\[
X_\ell = \sqrt{\alpha_{\ell}}X_0 + \sqrt{1 - \alpha_{\ell}}Z_\ell \quad \text{and} \quad X_{\ell+1} = \sqrt{\alpha_{\ell+1}}X_\ell + \sqrt{1 - \alpha_{\ell+1}}Z_{\ell+1},
\]

where both \( Z_\ell \) and \( Z_{\ell+1} \) are random variable with distribution \( \mathcal{N}(0, I) \). This results in

\[
X_{\ell+1} = \sqrt{\alpha_{\ell+1}}X_0 + \sqrt{1 - \alpha_{\ell+1}}Z_{\ell+1}
\]

for

\[
Z_{\ell+1} = \sqrt{\frac{\alpha_{\ell+1}(1 - \hat{\alpha}_{\ell})}{1 - \hat{\alpha}_{\ell+1}}} \hat{Z}_\ell + \sqrt{\frac{1 - \alpha_{\ell+1}}{1 - \hat{\alpha}_{\ell+1}}} Z_{\ell+1}.
\]
Therefore, we may take $\tilde{Z}_{t+1}$ as a reasonable approximation of $\tilde{Z}_t$, while sampling $\tilde{Z}_{t+1}$ is basically the same as sampling from $p_0(x) | x_{t+1}$. To summarize, we write

$$ q(x_0 | x_t, x_{t+1}) = q \left( \tilde{Z}_t = \frac{x_t - \alpha_t x_0}{\sqrt{1 - \alpha_t}} \bigg| x_t, x_{t+1} \right) $$

$$ \approx q \left( \tilde{Z}_{t+1} = \frac{x_t - \alpha_t x_0}{\sqrt{1 - \alpha_t}} \bigg| x_t, x_{t+1} \right) $$

$$ = q \left( X_0 = \frac{1}{\sqrt{\alpha_{t+1}}} \left( x_{t+1} - (x_t - \sqrt{\alpha} x_0) \sqrt{\frac{1 - \alpha_{t+1}}{1 - \alpha_t}} \right) \bigg| x_t, x_{t+1} \right) $$

$$ \approx p_0 \left( X_0 = \frac{1}{\sqrt{\alpha_{t+1}}} \left( x_{t+1} - (x_t - \sqrt{\alpha} x_0) \sqrt{\frac{1 - \alpha_{t+1}}{1 - \alpha_t}} \right) \bigg| x_{t+1} \right) $$

$$ = \mathcal{N} \left( \frac{\alpha_t (1 - \alpha_{t+1})}{\sqrt{\alpha_{t+1}} (1 - \alpha_t)} x_0 + \frac{x_{t+1}}{\sqrt{\alpha_{t+1}}} - \sqrt{\frac{1 - \alpha_{t+1}}{1 - \alpha_t}} x_t \right) $$

$$ h_\theta(x_{t+1}, \ell + 1), \text{diag}(\tau_{t+1}^2) \right) $$

$$ = \sqrt{\rho_\ell} \mathcal{N} \left( x_0 : \frac{1}{\sqrt{\alpha_t}}(x_t - \sqrt{1 - \alpha_t} \tilde{z}_{t+1}) , \rho_\ell \text{diag}(\tau_{t+1}^2) \right) , $$

where $\rho_\ell = \alpha_{t+1}(1 - \alpha_t)/(\alpha_t(1 - \alpha_{t+1}))$ and $\tilde{z}_{t+1}$ represents the noise predicted by the denoiser from $x_{t+1}$, that is,

$$ \tilde{z}_{t+1} = x_{t+1} - \sqrt{\alpha_{t+1}} h_\theta(x_{t+1}, \ell + 1) / \sqrt{1 - \alpha_{t+1}}. $$

Then, focusing on the arm that has been pulled at least once and ignoring the multiplicative constant that does not depend on $x_t$, we get

$$ q(y^g_0 | x_t) = \int q(y^g_0 | x_0) q(x_0 | x_t, x_{t+1}) dx_0 $$

$$ \approx \sqrt{\rho_\ell} \int q(y^g_0 | x_0) \mathcal{N} \left( x_0^a: \frac{1}{\sqrt{\alpha_t}}(x_t - \sqrt{1 - \alpha_t} \tilde{z}_{t+1}^a), \rho_\ell (\tau_{t+1}^a)^2 \right) dx_0 $$

$$ \propto \int \mathcal{N} \left( x_0^a : \frac{1}{\sqrt{\alpha_t}}(x_t - \sqrt{1 - \alpha_t} \tilde{z}_{t+1}^a), (\tau_{t+1}^a)^2 \right) dx_0 $$

$$ = \mathcal{N} \left( \hat{\mu}_{t+1}^a : \frac{1}{\sqrt{\alpha_t}}(x_t - \sqrt{1 - \alpha_t} \tilde{z}_{t+1}^a), (\sigma_{t+1}^a)^2 \right) $$

$$ \propto \mathcal{N} \left( x_0^a : \sqrt{\alpha_t} \hat{\mu}_{t+1}^a + \sqrt{1 - \alpha_t} \tilde{z}_{t+1}^a, \alpha_t (\sigma_{t+1}^a)^2 + \rho_\ell (\tau_{t+1}^a)^2 \right). $$

Subsequently, we can approximate the posterior distribution of $x_t$ using (5)

$$ q(x_t^a | x_{t+1}, y_0) \propto p_0(x_t^a | x_{t+1}) \mathcal{N} \left( x_t^a : \sqrt{\alpha_t} \hat{\mu}_{t+1}^a + \sqrt{1 - \alpha_t} \tilde{z}_{t+1}^a, \alpha_t (\sigma_{t+1}^a)^2 + \rho_\ell (\tau_{t+1}^a)^2 \right). $$

This results in the algorithm that we present in Section 3.2.

C Missing Experimental Details

In this section, we provide missing experimental details mainly concerning the used diffusion models and the construction of the problem instances. All the simulations are run on an Amazon p3.2xlarge instance equipped with an NVIDIA Tesla V100 GPU.

**Diffusion models.** In our experiments, we set the diffusion steps of the diffusion models to $L = 100$ and adopt a linear variance schedule from $1 - \alpha_1 = 10^{-2}$ to $1 - \alpha_{L} = 0.1$. Moreover, the models are trained to predict the clean sample $x_0$ instead of the noise $\tilde{z}_t$ since it is reported in [4] that this leads to better performance when the data are binary, and as explained below and shown in Figures 2 and 3, the mean rewards of the two considered tasks are nearly binary.

The denoiser itself is a 1-dimensional U-Net adapted from [9, 12]. We use 5 residual blocks, each block containing 6 residual channels. These numbers are rather arbitrary and do not seem to affect
much our results. More importantly, since the patterns that we want to learn do not have the spatial correlation that convolutional layers are designed for, we add a fully connected layer at the beginning to map the input to a vector of size $128 \times 6$, before reshaping these vectors into 6 channels and feeding them to the convolutional layers. In a similar fashion, we also replace the last layer of the architecture by a fully connected layer. We find out that these minimal modification already enable the model to perform well on our problems, but believe a thorough investigation into the architecture design would further benefit our approach.

These diffusion models are then trained with Adam for 15000 steps, with a learning rate of $5 \times 10^{-4}$, with a batch size fixed at 128 (which corresponds to 384 epochs).

**Construction of bandit instances.** We now give more details on how the mean reward vectors are constructed in the two problems. Some illustrations of the constructed instances and the vectors generated by the trained diffusion models are provided in Figures 2 and 3.

1. **Labeled Arms** ($K = 500$). As described in Section 4, in our first problem we are provided a set of 50 labels $L = \{1, ..., 50\}$. Each arm is associated to a subset $L^a$ of these labels with size $\text{card}(L^a) = 7$. To sample a new bandit task $B$, we randomly draw a set $L_B \subseteq L$ again with size 7. Then for each arm $a$, we set $\bar{\mu}^a = 1 - 1/4 \text{card}(L^a \cap L_B)$. Finally, to obtain the mean rewards $\mu$, we perturb the coordinates of $\bar{\mu}$ by independent Gaussian noises of standard deviation 0.1 and scale the resulting vector to the range $[0, 1]$.

2. **Popular and Niche** ($K = 200$). The arms are split into 40 groups of equal size. Conceptually, 20 of these groups represent the ‘popular’ items while the other 20 represent the ‘niche’ items. For each bandit task, we first construct a vector $\bar{\mu}$ whose coordinates’ values default to 0. However, we randomly choose 1 to 3 groups of niche items and the value of each of these items is set to 1 with probability 0.7 (independently across the selected items). Similarly, we randomly choose 15 to 17 groups of popular items and set their values to 0.8. Then, to construct the mean reward vector $\mu$, we perturb the values of $\bar{\mu}$ by independent Gaussian noises with standard deviation of 0.1. After that, we clip the values of the popular items to make them smaller than 0.95 and clip the entire vector to the range $[0, 1]$.

**UCB.** The most standard implementation of the UCB algorithm sets the upper confidence bound to

$$U^a_t = \hat{\mu}^a_t + \frac{\sigma'}{N^a_t} \sqrt{\log t}.$$  

(6)

Instead, in our experiments we use $U^a_t = \hat{\mu}^a_t + \frac{\sigma'}{\sqrt{N^a_t}}$. Eq. (6) is more conservative than our implementation. However, looking at Figure 1, we figure out that UCB without the log $t$ factor is already the most conservative algorithm in our experiments. Therefore, while in the long term UCB with (6) may achieve lower regret, with limited budget Thompson sampling with a wrong but good enough prior would generally be preferable [16].

**Regret.** For sake of completeness, we state here the definition of the (pseudo-)regret with respect to a sequence of arms $(a_t)_{t \in \{1, ..., T\}}$ in a single bandit task. Let $a^* \in \text{arg max}_{a \in A} \mu^a$ be an optimal arm. It is defined as $\text{Reg}_{T} = T \mu^{a^*} - \sum_{t=1}^{T} \mu^{a_t}$.
The mean reward vectors of 100 constructed bandit tasks.

The mean reward vectors of 80 constructed bandit tasks, grouped by labels and showing only 5 labels. Note that each arm has multiple labels and thus appears in multiple groups.

80 mean reward vectors generated by the trained diffusion model, grouped by labels and showing only 5 labels. Note that each arm has multiple labels and thus appears in multiple groups.

Fig. 2: Visualization of the mean reward vectors constructed in the Labeled Arms problem. Rows and columns correspond to tasks and arms. The darker the color the higher the value, with white and black representing respectively $\mu^a = 0$ and $\mu^a = 1$. While human eyes can barely recognize any pattern in the constructed vectors, diffusion models manage to learn the underlying patterns that become recognizable by humans only when the arms are grouped in a specific way.
(a) The mean reward vectors of 50 constructed bandit tasks.

(b) The mean reward vectors of 50 constructed bandit tasks. Reordered to put the arms of the same group together. The popular arms are on the right side of the figure.

(c) 50 mean reward vectors generated by the trained diffusion model. Reordered to put the arms of the same group together. The popular arms are on the right side of the figure.

**Fig. 3:** Visualization of the mean reward vectors constructed in the Popular and Niche problem. Rows and columns correspond to tasks and arms. The darker the color the higher the value, with white and black representing respectively $\mu^a = 0$ and $\mu^a = 1$. Diffusion models manage to learn the underlying patterns that become recognizable by humans only when the arms are grouped in a specific way.